



Dynamics in the international market segmentation of new product growth[☆]

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ABSTRACT

Prior international segmentation studies have been static in that they have identified segments that remain stable over time. This paper shows that country segments in new product growth are intrinsically dynamic. We propose a semiparametric hidden Markov model to dynamically segment countries based on the observed penetration pattern of new product categories. This methodology allows countries to switch between segments over the life cycle of the new product, with time-varying transition probabilities. Our approach is based on penalized splines and can thus be flexibly applied to any nonstationary phenomenon, beyond the new product growth context.

For the penetration of six new product categories in 79 countries, we recover the dynamic membership of each country to segments over the life cycle. Our findings reveal substantial dynamics in international market segmentation, especially at the beginning of the product life. Finally, we exploit the dynamic segments to predict the national penetration patterns of a new product before its launch and show that our forecasts outperform forecasts derived from alternate parametric and/or static methods. Our results should encourage multinational corporations to adopt dynamic segmentation methods rather than static methods.

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1. Introduction

Country segmentation is fundamental to any successful international marketing strategy (Steenkamp & Ter Hofstede, 2002). The globalization of firms and markets enhances the need for the cross-border exchange of experiences and market research that accounts for both similarities and differences across markets. With globalization comes an increased understanding of the similarities and differences between markets.

Various segmentation bases have been suggested for international markets; particularly relevant to the present paper is the segmentation of countries based on the sales, adoption or penetration patterns of new products over the life cycle (see, e.g., Gielens & Steenkamp, 2007; Helsen, Jedidi, & DeSarbo, 1993; Kumar, Ganesh, & Echambadi, 1998; Sood, James, & Tellis, 2009). Such segmentation is often used to select the sequence of countries to enter (Tellis, Stremersch, & Yin, 2003). Grouping countries based on the penetration patterns that new products show over time is especially relevant if one considers both the high financial stakes involved in introducing a new product globally and the

substantial differences that new product growth patterns show across countries and country segments (Dekimpe, Parker, & Sarvary, 2000; Desiraju, Nair, & Chintagunta, 2004; Gatignon, Eliashberg, & Robertson, 1989; Mahajan & Muller, 1994; Stremersch & Lemmens, 2009; Stremersch & Tellis, 2004; Van den Bulte & Stremersch, 2004; Van Everdingen, Fok, & Stremersch, 2009). The idea of such segmentation is that the penetration patterns of new products are likely to show similarities across countries when these countries face similar demand-side (e.g., national culture) and supply-side factors (e.g., regulation), as demonstrated by Stremersch and Lemmens (2009).

Prior research has cited multiple reasons why country segmentation on the basis of penetration patterns is relevant to companies. First, it enables cross-fertilization and experience sharing between managers of the same segment in different countries (Bijmolt, Paas, & Vermunt, 2004). Second, the sales evolution of a new product in one country can be used as reference point by managers in another country that belongs to the same segment (Steenkamp & Ter Hofstede, 2002). Third, international segmentation can improve forecasting accuracy regarding the growth of new products, especially prior to launch, which is similar in principle to analogical diffusion models (Bass, 2004; Ofek, 2005). A firm may also select a test country within a segment to explore the sales potential not only for that test market but also, by analogy, for the entire segment (Green, Frank, & Robinson, 1967). Firms can exploit the benefits of country segmentation both at the individual brand and at the product category level (e.g., brand diffusion versus category diffusion models or brand management versus category management).

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A key shortcoming of the country segmentation methods in the marketing literature is their static nature (Steenkamp & Ter Hofstede, 2002). An implicit assumption is made that the segments are stationary in both structure (segment composition or membership) and characteristics (segment profiles). In this model, the membership of a country to a segment does not vary over the product life cycle. For instance, Helsen et al. (1993) segment countries based on the time-invariant parameters of the Bass (1969) diffusion model in a mixture regression framework. Similarly, Jedidi, Krider, and Weinberg (1998) cluster movies according to their share-of-revenue patterns over time. Recently, Sood et al. (2009) propose a semiparametric model where they estimate diffusion curves, which they subsequently cluster using functional principal components (Ramsay & Silverman, 2005).

However, the nonstationary nature of new product adoption endangers the temporal stability of international segments, which are likely to show dynamics (Wedel & Kamakura, 2000). The combined studies of Tellis et al. (2003) and Stremersch and Tellis (2004) provide indirect evidence for the relevance of this time dependence. Using the same European diffusion data, they find that the factors that drive early growth (i.e., time-to-takeoff) are different from the factors that drive late growth (i.e., time-to-slowdown). Moreover, Golder and Tellis (2004) and Stremersch and Lemmens (2009) show that the influence of variables that affect new product growth varies over time.

In this paper, we demonstrate that country segmentation based on the penetration pattern of new product categories is not static but inherently dynamic. To accommodate dynamics, we develop a semiparametric hidden Markov model (HMM) that allows country segment membership to vary flexibly over time. The paper extends recent advances in time-varying household segmentation (e.g., Du & Kamakura, 2006; Paas, Vermunt, Bijmolt, & Journal of the Royal Statistical Society, 2007), dynamic customer value segmentation (Brangule-Vlagsma, Pieters, & Wedel, 2002; Homburg, Steiner, & Totzek, 2009) and customer relationship dynamics (Netzer, Lattin, & Srinivasan, 2008) to an international scope and combines them with recent advances in semiparametric modeling of new product growth patterns with penalized splines (Stremersch & Lemmens, 2009).

We apply our semiparametric dynamic segmentation method to country-level penetration data from six product categories of Information and Communication Technology (ICT) products and media devices across 79 developed and developing countries between 1977 and 2009. For this specific set of product categories, we identify three latent country segments that show substantial dynamics in membership probabilities and size over time, especially at the beginning of the product life cycle. Dynamic segments provide a better fit than static segments or segments based on geographic area (e.g., North America and western Europe). In addition, a semiparametric response function offers a better fit than a parametric specification. Our dynamic segmentation model also shows outstanding prelaunch forecasting performance, in most cases outperforming static and/or parametric segmentation methods. While we apply the model to product categories, it is also possible to apply it to brands because unlike (parametric) diffusion models such as the Bass diffusion model, our semiparametric approach does not impose a (behavioral) structure.

This paper has important implications for firms and international public policy bodies that use country segmentation methods. Many of these entities use an exogenously defined regional segmentation criterion (see e.g., Ghauri & Cateora, 2006, p. 492) or, if they are more sophisticated, a static model-based segmentation method. We show that both are inaccurate because segments are intrinsically dynamic. Therefore, both approaches may lead to inappropriate decision making and imprecise forecasts as compared with dynamic segments. We suggest that analysts change their current practice (i.e., static segmentation) and derive, for their respective industries and product categories, a dynamic segmentation of the countries in which they compete.

The remainder of the paper is organized as follows: the second section presents a short overview of the recent developments in

international segmentation modeling, the third section describes the methodological framework used to dynamically segment countries over time, the fourth section includes the data description and presents the empirical findings, and the final section presents managerial implications and conclusions.

2. Existing international segmentation methods

Most research has used established segmentation algorithms developed in the statistical literature, such as finite-mixture models (Helsen et al., 1993; Ter Hofstede, Steenkamp, & Wedel, 1999) or variations of k-means clustering (Chaturvedi, Carroll, Green, & Rotondo, 1997; Homburg, Jensen, & Krohmer, 2008; Kale, 1995). Few international segmentation studies have focused on the development of new methodological frameworks. Recently, scholars have proposed several new methods for studying international segmentation. In particular, hierarchical Bayesian models with segment-specific response parameters (Ter Hofstede, Wedel, & Steenkamp, 2002) allow spatial dependence within and between segments. The multilevel finite-mixture model proposed by Bijmolt et al. (2004) accounts for different levels of aggregation (e.g., consumer and country levels).

Other interesting ongoing methodological developments are functional data analysis and functional clustering, as in Sood et al. (2009) for product-country segmentation (or Foutz and Jank, 2010, for product segmentation). The main benefit of such approaches, beyond the flexibility that functional analysis offers as a nonparametric framework, is that the econometrician can cluster the growth curves of new products globally based on their functional shape. Any possible shape can be managed.

Time dependence remains an important concern in international segmentation. As discussed by Steenkamp and Ter Hofstede, “over time, the number of segments, segment sizes and structural properties of international segments may change. [...] This issue has not received rigorous attention yet.” (Steenkamp & Ter Hofstede, 2002, p. 209). From a managerial viewpoint, ignoring dynamics in international segments is likely to lead to suboptimal marketing strategies. From an estimation viewpoint, the violation of the assumption of stationarity may invalidate model estimation when the phenomenon under study is by nature nonstationary or when the data range spans a long time period, such as in diffusion studies. Recent methodological advances in segment dynamics modeling are (hidden) Markov models (Brangule-Vlagsma et al., 2002; Du & Kamakura, 2006; Homburg et al., 2009; Liechty, Pieters, & Wedel, 2003; Montgomery, Li, Srinivasan, & Liechty, 2004; Netzer et al., 2008; Paas et al., 2007; Ramaswamy, 1997). Applied to a collection of time series data, an HMM can identify, for each time period, the segment to which a realization belongs. HMMs allow segment membership to dynamically vary over time. Existing marketing applications have focused on customer or household segmentation and have modeled the finite-mixture response function in a parametric way. Our research extends the use of the HMM to international country segmentation and proposes the use of a semiparametric framework where time series data are not restricted to a specific functional form.

3. A semiparametric hidden Markov model for dynamic country segmentation

This section first describes the semiparametric hidden Markov model that we propose to dynamically segment countries. Then, we explain how we use this approach to make new product growth forecasts, and we present several alternative benchmark models.

3.1. A new dynamic segmentation framework

For every country i , with $i = 1, \dots, n$ and product category j , with $j = 1, \dots, J$, we observe a penetration pattern $y_{ij1}, y_{ij2}, \dots, y_{ijT_j}$, where T_{ij} is the number of sample points available for this product-country

combination. In our application, we define penetration in percentage as the number of devices or subscriptions used by a population divided by the number of users (see the data section). Prior to launch, we have $y_{ijt} = 0$ up to $t = 0$. Note that we consider duration time rather than calendar time because the goal of the analysis is to pool the penetration data of multiple product categories launched at different calendar times and to extract regularities or commonalities in diffusion patterns across countries. We let $\mathbf{y}_{it} = (y_{i1t}, \dots, y_{ijt}, \dots, y_{ijt})$ be the vector containing the penetration data in country i at time t of all product categories under consideration. As the number of observed sample points could differ per product category, the number of components in the vector \mathbf{y}_{it} is allowed to vary over time. We denote $T_i = \max_j T_{ij}$ and write, for notational convenience, for every country $T_i = T$.

We model the penetration of product j in country i at time t using a semiparametric hidden Markov model. A hidden Markov model is a probabilistic model of the joint probability of a collection of random variables, here $\{\mathbf{y}_{i1}, \dots, \mathbf{y}_{iT}\}$. The distribution of \mathbf{y}_{it} depends on the value taken by a hidden (or latent) state variable in the set $\{1, \dots, S\}$ of possible states, with S being the total number of hidden states. We denote (i, t) , the value of the hidden state variable for country i at time t . Countries that share the same value $s(i, t)$ belong to the same latent segment at time t . To allow for time-varying country membership to the latent segments, each country i follows a particular (hidden) sequence of states $s(i, 1), \dots, s(i, T)$, the state path. Using the time-varying segmentation basis \mathbf{y}_{it} , we estimate the model and obtain the probability of each country belonging to each latent segment at any given time point during the product life cycle.

3.1.1. Hidden Markov model

The model is composed of two parts: (i) a response component that connects the state variable to the observed responses at any given time point and (ii) a structural component that models changes in latent segments across time periods.

We model the penetration of product j in country i at time t , given that country i belongs to latent segment $s(i, t)$ as

$$y_{ijt} = f_{s(i,t)t} + g_{s(i,t)t} + \varepsilon_{ijt} \text{ with } \varepsilon_{ijt} \sim N(0, \sigma_t^2). \tag{1}$$

For every segment s , with $1 \leq s \leq S$, we denote that f_{st} is a segment-specific function, and g_{st} is the corresponding product-specific deviation from the segment function. Product-specific deviations vary between segments and capture the heterogeneity between products. These functions can be modeled in a parametric or semiparametric fashion. To keep our segmentation method as flexible as possible, we opt for a penalized spline semiparametric specification for both f_{st} and g_{st} , as detailed in the next subsection. Another option would have been to specify f_{st} and g_{st} as parametric functions of the time t (e.g., Bass, 1969). We compare both possibilities in the empirical analysis. Furthermore, we allow the error term in (1) to be heteroscedastic with $\sigma_t^2 = \sigma^2 t$ and each ε_{ijt} to follow a first-order autoregressive process with a common autoregressive parameter. This specification accounts for penetration curves being cumulative time series and provides a better fit to the data than a model with homoscedastic errors and/or without autocorrelation.¹ For a fixed time point, Eq. (1) can be interpreted as a finite-mixture model, defining a country segmentation based on the observed penetration values across product categories at that time period. Fig. 1 proposes a graphical representation of our semiparametric HMM in the case of three states or segments.

The structural component follows a first-order Markov chain. In particular, it assumes that membership to the latent segment at time t is affected only by segment membership at $t - 1$, but not by latent segment

membership at earlier periods. The initial latent segment probability $\pi_s = P(s(i, 1) = s)$ is the probability of belonging to segment s at $t = 1$ while the time-varying transition probability $\pi_{s_t, s_{t+1}}^t = P(s(i, t+1) = s_{t+1} | s(i, t) = s_t)$ denotes the probability of switching from segment s_t at t to segment s_{t+1} at $t + 1$, for $t = 1, \dots, T - 1$. The above probabilities are the same for all countries and are referred to as the prior probabilities. The prior probability that a country follows the state path s_1, \dots, s_T is then given by $\pi_{s_1} \pi_{s_1, s_2}^1 \dots \pi_{s_{T-1}, s_T}^{T-1}$ for any s_1, \dots, s_T , using the first-order Markov property.

In our approach, the transition probabilities are time-varying (or time-heterogeneous). In the nonstationary new product growth context, one can easily conceive that transition probabilities in the period immediately following the launch of a new product are likely to differ from transition probabilities after the product has matured. For instance, the first years after launch tend to show higher variability in state membership (lower stickiness of the states) than later years when the total market potential is almost reached. We allow these transition probabilities to vary freely over time. One could possibly extend this formulation by letting the probabilities depend on available (time-varying) covariates, as proposed by Paas et al. (2007). To identify the labels of segments, we use the restriction $f_{1t} < f_{2t} < \dots < f_{St}$ at each time point, allowing us to identify the segment with the largest value of the index s as the segment with the highest penetration level. A similar identification restriction was made in Netzer et al. (2008). Note that this restriction does not prevent a country from moving from the high- to the low-penetration segment (or vice versa), which would result in crossing penetration patterns.

Combining the semiparametric response model given by Eq. (1) and the first-order Markov model yields a semiparametric hidden Markov model. The total likelihood function is given by

$$\prod_{i=1}^n \sum_{s_1=1}^S \dots \sum_{s_T=1}^S \pi_{s_1} \pi_{s_1, s_2}^1 \dots \pi_{s_{T-1}, s_T}^{T-1} L(\mathbf{y}_{i1}, \dots, \mathbf{y}_{iT} | s_1, \dots, s_T), \tag{2}$$

with $L(\mathbf{y}_{i1}, \dots, \mathbf{y}_{iT} | s_1, \dots, s_T)$ being the conditional likelihood of the series for country i given the state path. This conditional likelihood is easy to compute. For instance, in the absence of serial correlation in the error terms in Eq. (1), we have $L(\mathbf{y}_{i1}, \dots, \mathbf{y}_{iT} | s_1, \dots, s_T) = \prod_{t=1}^T L_t(\mathbf{y}_{it} | s_t)$, where $L_t(\mathbf{y}_{it} | s_t)$ is the likelihood of a normal distribution with mean $\boldsymbol{\mu}_{st} = (\mu_{s1t}, \dots, \mu_{sjt})$ and covariance matrix R_t . Here, $\mu_{sjt} = f_{st} + g_{sjt}$ is the expected penetration level of product j in segment s at time t , and the matrix R_t is the covariance matrix of the error terms $\boldsymbol{\varepsilon}_{it} = (\varepsilon_{i1t}, \dots, \varepsilon_{ijt})'$. If we allow for autocorrelation in the error terms, the expression for the likelihood becomes slightly more complex because the distribution of \mathbf{y}_{it} depends on not only the state to which it belongs anymore but also $\mathbf{y}_{i, t-1}$ and the previous state.

3.1.2. Semiparametric modeling using penalized splines

To ensure full flexibility in the choice of the time-varying segmentation basis, we model the segment-specific function and the product-specific deviations in a semiparametric fashion, using penalized or p -splines. Previous diffusion research has argued that parametric diffusion models suffer from several limitations that semiparametric modeling can address. For instance, parameters tend to be biased when the observed time window is too short (Bemmaor & Lee, 2002; Van den Bulte & Lilien, 1997) or when data contain repeat purchases (Hardie, Fader, & Wisniewski, 1998; Van den Bulte & Stremersch, 2004). Penalized splines constitute a highly flexible and modular approach to model how a response variable is affected by covariates, in this case by duration time (Ruppert, Wand, & Carroll, 2003). Splines have become increasingly popular in medicine (e.g., Durban, Harezlak, Wand, & Carroll, 2005), finance (e.g. Jarrow, Ruppert, & Yu, 2004), and recently, in marketing (Kalyanam & Shively, 1998; Sloot, Fok, & Verhoef, 2006; Stremersch & Lemmens, 2009; Van Heerde, Leeflang, & Wittink, 2001; Wedel & Leeflang, 1998).

We can construct a spline as a linear combination of K linear bases, which are broken lines $(t - \kappa_k)_+$ truncated at knot κ_k with $0 \leq \kappa_k \leq T$

¹ Note that one could apply the logit transformation to the penetration data to ensure that the estimated penetration levels are between 0 and 1. In our particular application, all fitted values were in this interval.

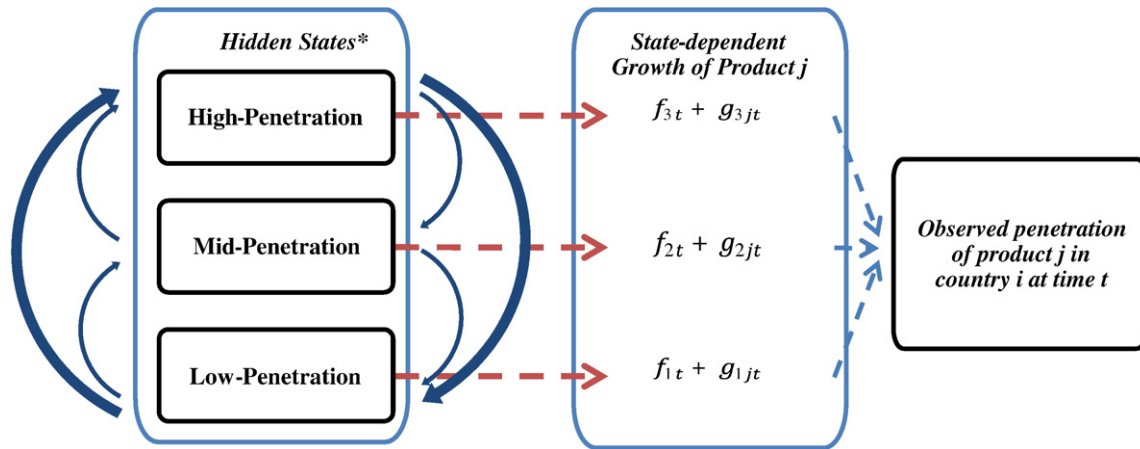


Fig. 1. A three-states hidden Markov model of new product growth.* Example of a 3-states hidden Markov model. The curved solid lines represent the transitions between states.

for $k = 1, \dots, K$ (i.e., the knot is the location where the broken lines are tied together).² If we combine such bases with different knots ranging from 0 to T and assign weights u_1, \dots, u_K to each, we can fit any nonlinear, smoothed curve $h(t) = \sum_{k=1}^K u_k(t - \kappa_k)_+$. To ensure the smoothness of the curve, these weights are *penalized* (i.e., they are subject to the constraint $\sum_{k=1}^K u_k^2 < U$, for some constant U) (see Ruppert et al., 2003, pp. 65–67, for more details). This weighting mechanism explains why splines are called penalized splines. The number of knots, K , must be large enough to ensure the flexibility of the curve. The level of smoothing of the penalized splines is controlled by the variance of u_k (i.e., σ_u^2). A large variance corresponds to a wiggly function while a small σ_u^2 yields a smooth function. Note that the variance is estimated using the maximum likelihood (Wand, 2003).

The segment-specific function (see Eq. (1)) is written as a penalized spline

$$f_{st} = \beta_s t + \sum_{k=1}^K u_{sk}(t - \kappa_k)_+ \quad (3)$$

with fixed slope parameters β_s and random coefficients $u_{sk} \sim N(0, \sigma_u^2)$. It can be interpreted as the average penetration pattern observed in the country segment s . This component reflects the regularities in the penetration patterns that new products exhibit in this specific country segment. In turn, deviations for product j from the segment-specific function are modeled similarly as

$$g_{sjt} = \alpha_{sj} t + \sum_{k=1}^K v_{sjk}(t - \kappa_k)_+ \quad (4)$$

with the random parameters $\alpha_{sj} \sim N(0, \sigma_\alpha^2)$ and $v_{sjk} \sim N(0, \sigma_v^2)$.

The above formulation treats products as being nested into segments and therefore accounts for product-segment interaction effects. It allows the product-specific deviations to vary across country segments. Controlling for the product deviations allows us to define an international segmentation solution that is generalizable to the set of product categories considered, making an abstraction of the product-specific peculiarities and the noninformative variation in the data (e.g., measurement error). Note that the resulting country segments may be different if one considers a different pool of product categories (e.g., high tech products versus kitchen and laundry appliances). For firms, the best approach would be to consider their product divisions and pool across all the products of each respective division.

3.1.3. Parameter estimation

The model parameters are estimated using a maximum likelihood and computed using the Expectation-Maximization (EM) algorithm.

² The notation $h(x) = (x)_+$ indicates that the function h equals zero when $x < 0$ and equals x for $x \geq 0$.

We refer to Zucchini and MacDonald (2009) for details on the EM algorithm applied to the hidden Markov model.³ The outline follows: denote θ as the vector collecting all unknown parameters of the model, including the initial state probabilities, the transition probabilities, and the unknown f_{st} and g_{sjt} ; and assume that, at the k^{th} step of the algorithm, an estimate θ_k is available. In addition, P_k and L_k are the probability and the likelihood, assuming that $\theta = \theta_k$. In the E-step, we compute the following: (i) the *posterior segment membership probabilities*, $P_k(s(i, t) = s | \mathbf{y}_{i1}, \dots, \mathbf{y}_{iT})$, (i.e., the probability of belonging to segment s at time t for country i , conditional on the observed time series) and (ii) the *posterior segment transition probabilities*. These posterior probabilities can be computed efficiently using the Baum–Welch forward-backward algorithm (Baum, Petrie, Soules, & Weiss, 1970, see also Paas et al., 2007). By averaging the posterior probabilities $P_k(s(i, 1) = s | \mathbf{y}_{i1}, \dots, \mathbf{y}_{iT})$ over all the countries, we obtain an update of the estimates of the initial state probabilities π_s for $s = 1, \dots, S$. Similarly, by averaging the posterior segment transition probabilities over all countries, we obtain an update of the transition probabilities $\pi_{s_1 s_2}^t$. In the M-step, we maximize the expected log-likelihood, leading us to maximize

$$\sum_{i=1}^n \sum_{s_1=1}^S \dots \sum_{s_T=1}^S P_k(s(i, 1) = s_1, \dots, s(i, T) = s_T | \mathbf{y}_{i1}, \dots, \mathbf{y}_{iT}) \log L_k(\mathbf{y}_{i1}, \dots, \mathbf{y}_{iT} | s_1, \dots, s_T) \quad (5)$$

with respect to the unknown parameters f_{st} and g_{sjt} appearing in the likelihood, yielding a new estimate θ_{k+1} . Maximizing the expected log-likelihood corresponds to computing a weighted maximum likelihood estimator. Alternatively, the observations could have been randomly assigned to the latent states according to the posterior probabilities of the state paths. The standard estimators are then computed from the data allocated to the different states. This approach is called the Stochastic Expectation-Maximization (SEM) algorithm and was first proposed by Celeux and Govaert (1991). We then iterate the E and M steps until we reach convergence of the model fit criteria. We start the algorithm with equal posterior latent segment membership probabilities $1/S$, and equal latent posterior transition probabilities $1/S$, for $t = 1, \dots, T - 1$.

Note that our model-based segmentation provides richer information than “hard” clustering algorithms (e.g., k -means), as it yields the probabilities of each country belonging to each of the various segments at every time point.

³ Most textbooks only present the homogenous version of the HMM, where the transition probabilities are not time-dependent. We carefully adapted all formulas – in particular of the Baum–Welch algorithm – to the nonhomogeneous case. For brevity, we only report here the most important ingredients of the EM algorithm. Full details are available from the first author upon request.

3.1.4. Model fit criteria and the optimal number of segments

As is commonly done in segmentation analysis (see e.g., Wedel & Kamakura, 2000), we evaluate the fit of the model by computing the Bayesian Information Criterion (BIC) based both on the estimated likelihood given in Eq. (2) and on the total number of free parameters in the model (Greene, 2003, p. 160). The latter includes the transition probabilities and starting probabilities and the parameters in the specification of both the segment curve (Eq. 3) and the product-deviation curve (Eq. 4). In the context of our semiparametric mixed-effect model, we compute the number of free parameters following the approach proposed by Vaida and Blanchard (2005, see also Ruppert et al., 2003), yielding a number ranging between the sum of the number of fixed effects and variance components and the sum of the number of fixed and random effects.

In addition to the Bayesian information criterion, we also evaluate the separability of the segments and the stability of the segmentation to changes in the data (i.e., its robustness). The normalized entropy criterion (NEC) can be computed to investigate the degree of separation in the posterior probabilities (Grover & Vriens, 2006, pp. 402–403, 416). A lower NEC indicates that the segments are separated well from one other. In addition, a value less than one indicates that the identified segmentation structure does exist (Biemacki, Celeux, & Govaert, 1999).

To study the stability of the segmentation to changes in the data, we apply the model explorer algorithm of Ben-Hur, Elisseeff, and Guyon (2002), which makes use of the following cross-validation. We randomly separate 10% of the countries and apply the model to the remaining 90%. Repeating this operation 10 times, we obtain 10 different segmentations, in which pairwise similarity indices are computed using the popular Adjusted Rand Index (ARI) generalized to probabilistic segmentation (Anderson, Bezdek, Popescu, & Keller, 2010; Hubert & Arabie, 1985). The ARI takes values between -1 and 1 , where 0 indicates that the clustering is due to chance, and 1 indicates a perfect similarity. A high average similarity index suggests a high stability of the segmentation.

To select the optimal number of segments, we compute these statistics for different numbers of segments S , and the model yielding the most satisfactory fit, separability and stability is selected.

3.2. Forecasting procedure

Country segmentation can be a powerful instrument for prelaunch forecasts. Prior to the first launch of a product, firms do not observe any product-specific information on the actual adoption of the new product or product category. In such cases, firms can rely on test market data, consumer surveys or “clinics” (Blattberg & Golanty, 1978; Urban, Hauser, & Roberts, 1990), advance purchase orders (Moe & Fader, 2002), and/or the available sales or adoption history of similar products introduced in the past using analogical diffusion models (Lee, Boatwright, & Kamakura, 2003; Ofek, 2005). For international markets, firms can also pool the information on similar products across multiple countries, rather than using the information for a single country (Talukdar, Sudhir, & Ainslie, 2002). In this context, country segments should indicate the relevant set of countries to be considered in prelaunch forecasting. If the premise that we stated above is correct (i.e., that information contained in dynamic segments of countries is richer than information contained either in static segments of countries or in single countries), then the dynamic segmentation method should outperform alternative methods based on static segments of countries or past patterns in single countries. This idea is similar to the spatial model proposed by Bronnenberg and Sismeiro (2002) that infers data in markets where little or no data are available using information available from other geographic locations.

The use of segments to make prelaunch forecasts depends on the following: (i) whether forecasts are made before a new product (category) is initially launched (i.e., is not available in any country yet) or (ii) whether forecasts are made before a new product (category) is

launched in a given country (*local launch*) but has already been launched in other countries.

3.2.1. Forecasts before the first international launch

Before a new product (category) j_0 is launched for the first time, we predict its penetration in country i at duration time t (i.e., t years after launch) as the value at t of the segment-specific curve to which country i belongs. Forecasts can be obtained for all prediction horizons in this fashion. In the case of dynamic segments, the set of similar countries used to build forecasts changes over the product life cycle as segment membership becomes time-varying. If the segmentation is probabilistic, as it is for our proposed HMM, the predicted curve for the focal country i refers to a weighted average of all segment-specific curves. That is,

$$\hat{y}_{ij_0t} = \sum_{s=1}^S w_{ist} \hat{f}_{st}, \quad (6)$$

where \hat{f}_{st} is the estimate of the segment-specific curve for segment s . The weights w_{ist} are the posterior probabilities that the focal country i belongs to the given segment at the corresponding time t ; so $w_{ist} = P(s(i, t) = s | \mathbf{y}_{i1}, \dots, \mathbf{y}_{iT})$ results from the EM algorithm.

3.2.2. Forecasts before local launch

Companies often make use of waterfall entry strategies by spreading national introductions over a given time span. In this context, it is possible to use experience from previously entered markets to improve our forecasts. Specifically, in cases where the focal country i is not the first international entry and the product (category) j_0 has been introduced in similar countries (i.e., in countries belonging to the same segment), we make forecasts in focal country i using the penetration data of the other member countries of the segment. Again, with dynamic segments, the set of similar countries used to build forecasts changes over the product life cycle. We denote \bar{y}_{sj_0t} as the average penetration level in all previously entered countries that belongs to segment s for product j_0 , and for which data are available at time t . We predict the penetration of product j_0 in country i at time t as the following weighted average:

$$\hat{y}_{ij_0t} = \sum_{s=1}^S w_{ist} \bar{y}_{sj_0t}. \quad (7)$$

where the weights are defined as in Eq. (6).

3.3. Benchmarks

We evaluate the fit and forecasting performance of our semiparametric hidden Markov model for dynamic segmentation against a number of benchmarks. Benchmarks are chosen to allow us to assess the contribution of each of the following dimensions characterizing our approach:

- (i) *Semiparametric vs. parametric response model;*
- (ii) *Multi-country vs. single-country segments;*
- (iii) *Model-based vs. a priori-defined segmentation;*
- (iv) *Dynamic vs. static segments.*

The various benchmarks are listed in Table 1 and are subsequently described in turn.

3.3.1. Semiparametric vs. parametric response model

First, we assess whether our semiparametric spline-based HMM yields a better fit and forecasting performance than a parametric variant. To do so, we replace the segment-specific and product-deviation response functions in Eq. (1) with parametric equivalents. We use the Bass (1969) mixed-influence model. For completeness, we also implement each of the methods below in a parametric and semiparametric

Table 1
Model comparison.

Models	Semiparametric vs. parametric response model	Multi-country vs. single-country segments	Model-based vs. a priori-defined segmentation	Dynamic vs. static segments
<i>Benchmarks</i>				
Single-country segments	Parametric	Single-country	A priori	Static
	Semiparametric	Single-country	A priori	Static
A priori-defined segments (geographic regions)	Parametric	Multi-country	A priori	Static
	Semiparametric	Multi-country	A priori	Static
Static segments	Parametric	Multi-country	Model-based	Static
	Semiparametric	Multi-country	Model-based	Static
Dynamic segments	Parametric	Multi-country	Model-based	Dynamic
<i>Our proposal</i>				
Dynamic segments	Semiparametric	Multi-country	Model-based	Dynamic

way, allowing us to assess the systematic contribution of the p -splines approach through all the approaches.

3.3.2. Multi-country vs. single-country segments

Second, we assess whether grouping countries into segments (multi-country segments) yields a better fit and forecasting performance than considering each country as a separate segment (single-country segments). A so-called *single-country segments* model with one country per segment can be represented by Eqs. (3) and (4), where the segment-specific parameters are replaced by country-specific parameters. A parametric and a semiparametric version are considered. Note that the parametric version is equivalent to the multi-product, multi-country, Bass model proposed by Talukdar et al. (2002) without covariates. The resulting segmentations are static, as the segmentation membership is constant over time.

3.3.3. Model-based vs. a priori-defined segmentation

Third, we assess whether model-based segmentation yields a better fit and forecasting performance than a-priori segmentation. To do so, we replace the segment-specific parameters in Eqs. (3) and (4) with geographic region indices, yielding *a priori-defined segments*. We follow the geographic classification established by the United Nations Statistics Division: Africa, southeast Asia, eastern Europe, Latin America, the Middle East (Western, Central and Southern Asia), North America, Oceania, and western Europe. Table 2 depicts a list of all countries in our study by geographic region. Segments are static as the segmentation membership is constant over time.

3.3.4. Dynamic vs. static segments

Fourth, we assess the relevance of allowing for segment dynamics by comparing the proposed hidden Markov models to finite-mixture

models, further denoted as *static segments*. Countries were not allowed to change segment membership over time. We implement both a parametric finite-mixture Bass model, as proposed by Helsen et al. (1993), and a semiparametric finite-mixture splines model, along the same lines as the functional clustering approach suggested by James and Sugar (2003).

Note that all the models are implemented within the same framework. The parametric models are obtained by replacing the segment (and product) curves in Eq. (1) with the Bass diffusion function, depending on three segment- and product-specific parameters. The *single-country segments* approach corresponds to $S = N$ segments (i.e., each country is a single segment). The a-priori segmentation approach replaces the probabilities in Eq. (2) with an indicator of the known cluster membership. Finally, the finite-mixture model is a special case of the HMM, where the transition matrix is diagonal. In total, 8 different models are obtained that represent all possible cases. Note that we do not need a full factorial design of 2^4 combinations because the 8 omitted combinations of factors are impossible cases (e.g., single-country segments are, by nature, neither model-based segments nor dynamic, and a priori-defined segments are also not dynamic). When the segmentation is model-based, the number of segments, as reported in the second column of Table 5, is selected according to the BIC criterion.

4. Data

We gathered annual data on the percentage penetration of six new product categories among households in 79 countries. The data source is Euromonitor. The new product categories are ICT products and media devices, and we therefore expect them to exhibit similarities in their penetration patterns. The products include CD players, DVD players (including DVD recorders), home computers, Internet subscriptions,

Table 2
Included countries by geographic region.

Africa	Asia	Eastern Europe	Latin America	Middle East	North America	Oceania	Western Europe
Algeria	China	Belarus	Argentina	Azerbaijan	Canada	Australia	Austria
Cameroon	Hong Kong	Bosnia Herz.	Bolivia	Bahrain	USA	New Zealand	Belgium
Egypt	India	Bulgaria	Brazil	Iran			Denmark
Morocco	Indonesia	Croatia	Chile	Israel			Finland
Nigeria	Japan	Czech Rep.	Colombia	Jordan			France
South Africa	Malaysia	Estonia	Costa Rica	Kazakhstan			Germany
Tunisia	Philippines	Georgia	Dom. Rep.	Kuwait			Greece
	South Korea	Hungary	Mexico	Pakistan			Ireland
	Taiwan	Latvia	Peru	Qatar			Italy
	Thailand	Lithuania	Uruguay	Saudi Arabia			Netherlands
	Vietnam	Macedonia	Venezuela	Turkey			Norway
		Poland		Turkmenistan			Portugal
		Romania		UAE			Spain
		Russia					Sweden
		Slovakia					Switzerland
		Slovenia					United Kingdom
		Ukraine					

mobile phones, and cable television. As we define penetration as the number of devices or subscriptions used by a population divided by the number of users (member households or individuals), penetration could exceed 100% (i.e., when users own multiple devices) and could decrease over time (i.e., when users disadopted products). The set of 79 countries (see Table 2) is global, consisting of western and eastern European countries, North American and Latin American countries, and African and Middle Eastern countries, and thus contains both developing and developed countries, as recommended by Burgess and Steenkamp (2006).

The database covers the period from 1977 to 2009. Because the various technologies are introduced at different times during this period, the starting date of each series differed across product categories and countries. Note that we observe a maximum of 25 years of data per product-country combination. Several technologies had presumably not reached half of their market potential in 2009, as there is no inflection point within the data range. In total, we obtain data on 398 product-country combinations. We achieve full country coverage for DVD players, Internet subscriptions and home computers while some countries are missing data for CD players, mobile phones and cable television.

5. Results

In this section, we first present the results of the dynamic segmentation model applied to the aforementioned set of new product categories. Next, we demonstrate the superior fit and forecasting accuracy of the dynamic segmentation method as compared with the benchmark methods introduced above.

5.1. Dynamic country segments in new product growth

We estimate the semiparametric hidden Markov model for various numbers of segments. To determine the appropriate number of segments, we compute various model performance statistics (see the methodology section), including (i) the overall fit using the Bayesian information criterion (BIC), (ii) the separability of the segments using the normalized entropy criterion (NEC) and (iii) the stability of the segmentation solution using the adjusted Rand index (ARI). All results are reported in Table 3. The lowest BIC and NEC are obtained using a 3-segment solution. For the ARI, the 2-segment and 3-segment solutions yield the highest stability to changes in the data. Therefore, we opt for the 3-segment solution.

Fig. 2 shows the segment-specific penetration pattern for these three latent segments. We subsequently label the segments according to the level of the dependent variable as “low-penetration” (segment 1), “mid-penetration” (segment 2) and “high-penetration” (segment 3) segments. The low-penetration segment exhibits slow growth over the complete time range, with penetration levels reaching about 40% of the households 25 years after introduction. The mid-penetration segment shows a higher growth rate than the low-penetration segment, especially when the product category has been on the market for more than 8 years. Finally, the high-penetration segment shows a substantially faster and higher diffusion rate over the complete time range

Table 3
Model fit, segment separability and stability for various number of segments.

Number of segments (S)	Bayesian Information Criterion (BIC)	Normalized Entropy Criterion (NEC)	Adjusted Rand Index (ARI)
1	41,818	1.00	–
2	38,111	0.43	0.31
3	36,187	0.25	0.29
4	39,441	0.58	0.07
5	46,049	5.62	0.02
6	48,127	5.17	0.04

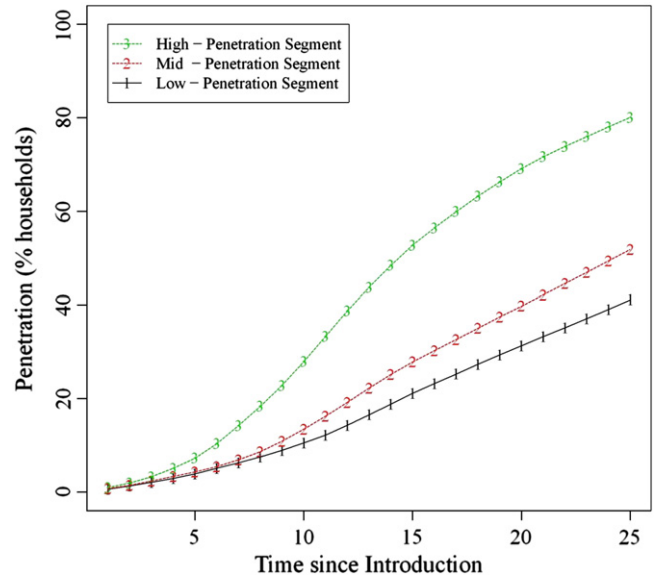


Fig. 2. Segment-specific penetration patterns of the high-, mid- and low-penetration segments.

than the other two segments. Penetration in the high-penetration country segment accounts for 80% of the households after 25 years.

To assess the reliability of the estimates of the segment-specific functions, we use the parametric bootstrap procedure described in Zucchini and MacDonald (2009, p. 55). The bootstrap procedure captures the uncertainty in both the segment membership and the estimation of the segment-specific curves. We obtain an estimate of the covariance matrix of the estimator and compute the standard error of the difference between each pair of estimated curves. The resulting t-statistic values over the product life cycle (see the Appendix) confirms that the high-penetration segment exhibits consistently higher penetration levels than the other segments over the complete time range while the low- and mid-penetration segment curves only became significantly different from each other once 8 years has passed since the introduction of the product. At the early stages of diffusion, the information in the data is still too scarce to be able to clearly distinguish segments; penetration is low in both segments. These findings are presented in Table 4, where the transition probabilities between these two segments are greater than 40% during the first 5 years after launch. In time, these two segments become more distinct. This phenomenon is also shown in Table 4, where we can see that the transition probabilities decrease between 6 and 15 years after introduction.

Table 4
Average transition probabilities over 3 successive time ranges (and standard deviations within brackets).

Period t - 1	Period t		
	Segment 1	Segment 2	Segment 3
<i>From t = 1 to 5 years after introduction:</i>			
Segment 1	52.06% [0.08]	41.70% [0.03]	6.24% [0.06]
Segment 2	40.89% [0.04]	54.12% [0.08]	4.98% [0.07]
Segment 3	1.68% [0.03]	1.66% [0.03]	96.66% [0.06]
<i>From t = 6 to 15 years after introduction:</i>			
Segment 1	91.42% [0.18]	8.55% [0.18]	0.03% [0.00]
Segment 2	8.94% [0.18]	91.05% [0.19]	0.02% [0.00]
Segment 3	0.00% [0.00]	0.00% [0.00]	100.00% [0.00]
<i>From t = 16 to T years after introduction:</i>			
Segment 1	82.63% [0.04]	17.37% [0.04]	0.00% [0.00]
Segment 2	16.79% [0.05]	83.21% [0.05]	0.00% [0.00]
Segment 3	0.00% [0.00]	0.00% [0.00]	100.00% [0.00]

Specific to our dynamic segmentation approach, we observe countries changing segments over the life cycle of the product. These changes are governed by the time-varying transition probability matrices. Table 4 reports the average transition probabilities (in percentages) between segments during three consecutive time spans: year 1 to year 5, year 6 to year 15, and year 16 to year 25. The matrices provide information on the *stickiness* in each segment over time. In particular, diagonal elements indicate how likely countries are to remain in the same segment over the product life cycle. In contrast, the nondiagonal elements capture the existing dynamics in segment membership. We find that the dynamic nature of the international market segmentation of a new product is most pronounced at the beginning of the product life cycle. When product categories became more mature, the dynamic nature of segment membership decreased. Compared with the low- and mid-penetration segments, the high-penetration segment shows little dynamics. Most segment changes occur between the low- and mid-penetration segments, and most changes occur between neighboring segments (e.g., between segments 1 and 2 but rarely between segments 1 and 3).

As countries switch segments over the product life cycle, segment membership probabilities evolve, as depicted by Fig. 3. Fig. 3 reports the evolution of the prior membership probabilities among the three segments. Interestingly, we find that the high-penetration segment (segment 3) is small in terms of the number of members during the first years of the product life cycle and gradually increases in size as the product categories mature. In contrast, the low-penetration segment is initially the largest segment and gradually loses members over time. Finally, the size of the mid-penetration segment is much less variable. For mature product categories, country segments display similar sizes.

Beyond the prior segment membership probabilities, we also estimate the time-varying posterior segment membership probabilities per country for each country in our data (i.e., the probability of each country i belonging to segment s at time t). Fig. 4 reports these membership probabilities among 6 leading industrial nations: the USA, the United Kingdom, Japan, Germany, France and Italy. Fig. 5 reports these probabilities among 6 leading emerging markets, often referred to as the BRICS + M countries: Brazil, Russia, India, China, South Africa and Mexico. Collectively, these 12 countries capture more than half of the global population and economic activity.

Figs. 4 and 5 clearly illustrate how countries stochastically switch between segments over time and allow the reader to visualize the most likely state path of each country. In general, we observe three different types of paths in our data set. Some countries belong with a very

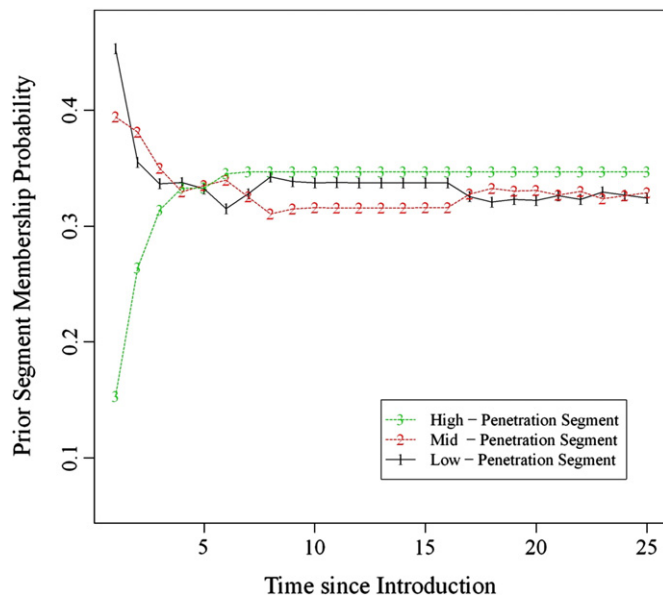


Fig. 3. Evolution of the prior segment membership probabilities.

high probability to the high-penetration segment over the complete time span (except perhaps for the first few years after product introduction) and therefore display little membership dynamics over time. This is true for the 6 leading industrial nations in Fig. 4 and also for most western European countries (Austria, Belgium, Finland, Ireland, the Netherlands, Norway, Portugal, Spain, Sweden and Switzerland); Canada; Australia and New Zealand; and the most developed southeast Asian and Middle East nations (Bahrain, Hong Kong, Israel, Kuwait, South Korea, Taiwan, and the United Arab Emirates).

Another set of countries tend to switch between the low- and mid-penetration segments during the first 6–7 years after product introduction and subsequently settle in the mid-penetration segment in later periods. These countries include Russia, China and South Africa in Fig. 5, and they also include many eastern European countries (Bosnia-Herzegovina, Bulgaria, the Czech Republic, Estonia, Hungary, Lithuania, and Slovenia) and several other nations from the Middle East, southeast Asia (Kazakhstan, Qatar, Saudi Arabia) and Latin America (Columbia and Costa Rica).

A final set of countries show comparable probabilities to belong to (and therefore tend to switch between) the low- and mid-penetration segments during the first 6–7 years after product introduction but then tend to subsequently remain in the low-penetration segment. These countries include Brazil, India and Mexico in Fig. 5; several less-developed or developing economies from various parts of the world, mostly Latin American countries (Argentina, Chile, the Dominican Republic, Peru and Venezuela) and many African countries (Algeria, Cameroon, Egypt, Morocco, Nigeria, and Tunisia); some eastern European countries (Belarus, Croatia, Georgia, Poland, Romania, Slovakia and Ukraine), and Asian (Indonesia, Philippines and Vietnam) and Middle Eastern countries (Iran, Jordan, Pakistan, Turkey and Turkmenistan). Our results are in line with the findings of Dekimpe et al. (2000), Van den Bulte and Stremersch (2004), and Van Everdingen et al. (2009), among others, who find the adoption timing of new products to be negatively correlated with Gross Domestic Product.

To conclude the investigation of our results, we compare the fit of our dynamic segmentation model with the alternative specifications described in section 3.3 and summarized in Table 1. In Table 5, we report the Bayesian Information Criterion (BIC), the log-likelihood and the number of segments chosen for each model. First, our results show that a semiparametric specification leads to a better fit performance than its parametric counterpart in terms of BIC. The flexibility of semiparametric models paid off. The only exception is for the *single-country segments* approach because the BIC penalizes for the high number of country-specific parameters involved in the semiparametric model. However, the fit performance of the *single-country segments* models remains inferior to the performance of all *multi-country segments* models.

Second, our results clearly indicate that the model-based segmentation approaches (i.e., the static and dynamic segments) outperform in fit the a priori-defined regional segments, both for the parametric and semiparametric versions. Such findings support the current knowledge in marketing research that domain-based segmentation bases should be favored over general segmentation bases (Steenkamp & Ter Hofstede, 2002).

Third, the fit comparison also shows that the semiparametric dynamic segmentation approach yields the best fit performance of all specifications.

To conclude, our proposed semiparametric dynamic segmentation approach demonstrates better in-sample fit than the other models. In the next section, we will assess the hold-out predictive performance of our approach.

5.2. Prelaunch predictive performance

We assess the relative predictive performance of the dynamic segmentation model against the alternative approaches described in section 3.3. The prediction task consists of forecasting the future

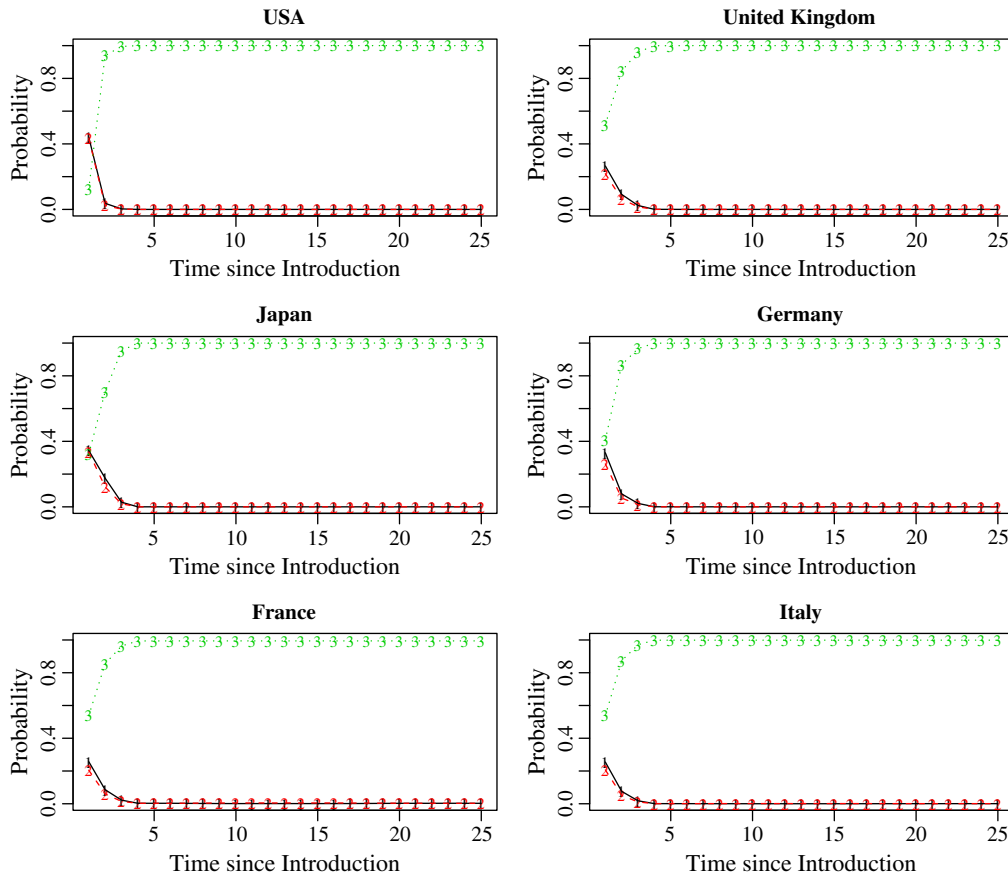


Fig. 4. Dynamic segment membership probability path for 6 leading industrial nations: USA, United Kingdom, Japan, Germany, France and Italy. Labels indicate (1) the low-penetration segment, (2) the mid-penetration segment and (3) the high-penetration segment.

penetration levels reached by a new product category in each country before its launch time in this country. All forecasts are made prior to the product introduction and for various prediction horizons (further denoted h), where h ranges from 1 year to 5 years after launch (i.e., 5 years-ahead forecasts).

For forecasts made before the first international launch, no information on the actual adoption of the new product is available yet. If we denote \tilde{t}_{ij_0} as the year when a new product j_0 is introduced in country i , the estimation sample to forecast penetration of product j_0 in country i at prediction horizon h only includes the penetration data of other products available prior to \tilde{t}_{ij_0} . In other words, we cut our data sample according to the calendar time at \tilde{t}_{ij_0} . All data corresponding to the years after product j_0 has been launched in country i do not belong in the estimation sample.

For subsequent entries (all countries entered after the first country entered), some information on the actual adoption of the product became available. More specifically, the estimation sample to predict the penetration of a new product j_0 in a subsequent entry i' at prediction horizon h also includes the penetration data on product j_0 in previously-entered countries up to the introduction time in the focal country i' . We use the penetration of product j_0 in the focal country for all available years as a hold-out sample. Thus, for each product-country combination, we construct a different estimation and hold-out sample, divided according to calendar time. This framework replicates the data context practitioners face when making pre-launch forecasts.

Because the goal is to use the penetration data of older product categories to forecast the penetration of new product categories, we focus in the analysis on the most recent product categories in our data set, DVD players and Internet, which were both introduced in the 1990s. We assess the predictive accuracy of each method by

computing the mean absolute deviation (MAD) between the predicted value and the actual value across all countries per prediction horizon. Table 6 reports the average MAD of the prelaunch forecasts for all methods per product category over various prediction horizons, from $h = 1$ to $h = 5$. In practical terms, the MAD can be interpreted as the average absolute deviation from the actual penetration level. For instance, a $MAD = 1.00$ indicates that our forecasts deviate from the actual penetration level by 1.00 unit. If the actual penetration is 20% of households, the corresponding forecast averages between 19% and 21% of households. The best method is reported in bold for each product category. Note that in Table 6, the results for the forecasts made before the first international launch and before a local launch are lumped together to compute the MAD.

The out-of-sample forecasting comparison supports four main conclusions. First, semiparametric models give more accurate forecasts than parametric models based on the Bass specification. Second, relying on multi-country segments rather than single countries generally improves prelaunch forecasts, indicating that segmentation allows for the identification of similar countries and can help the analyst decide how to account for data available from other countries. Information is thus gained by pooling across countries. For DVD players, forecasts made at the multi-country segment level mostly outperform forecasts made using country-specific information from one country only. For Internet subscriptions, this is also the case for the semiparametric models while all parametric models perform quite poorly. Third, we do not find substantial differences in forecasting performance between the a priori-defined segments and the static model-based segments. Fourth, dynamic segments yield better forecasts than static segments, confirming that there are substantial dynamics in country segment membership, which the static segmentation approach do not account for.

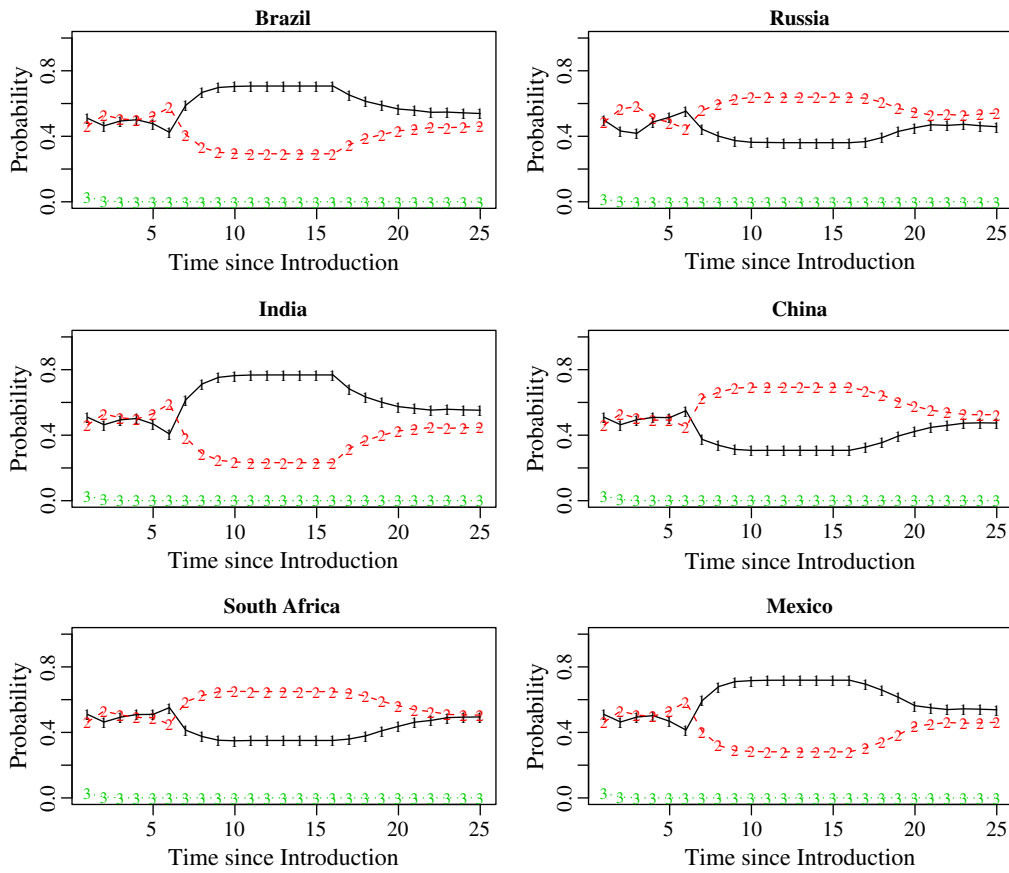


Fig. 5. Dynamic Segment Membership Probability Path for 6 Leading Emerging Markets: Brazil, Russia, India, China, South Africa and Mexico. Labels indicate (1) the low-penetration segment, (2) the mid-penetration segment and (3) the high-penetration segment.

In practical terms, our dynamic segmentation method yields low average absolute forecast deviations both for DVD players and Internet subscriptions. For the former, we find that our forecasts deviate by only .64 units (i.e., 0.64% of the households in the country) from the actual penetration levels 5 years after introduction. For Internet subscriptions, the dynamic segmentation yields forecasts that deviate by 2.22 units (i.e., 2.2% of the households in the country) on average from the actual penetration level 5 years after introduction.

In sum, the semiparametric dynamic segmentation approach offers the lowest forecast errors for all prediction horizons for both product categories, suggesting that country segments show substantial dynamics and that the flexible semiparametric specification of the penetration pattern is appropriate. These results are particularly encouraging given that the forecast of the first five years after launch

is probably the most crucial for managers when they plan to launch a new product on the market.

6. Conclusions and discussion

Recent calls in the marketing literature have highlighted the need for a dynamic modeling framework when approaching nonstationary marketing phenomena (e.g., Lemmens, Croux, & Dekimpe, 2007; Pauwels et al., 2004), such as new product adoption. Furthermore, in the country segmentation literature, Steenkamp and Ter Hofstede (2002) have cautioned that the static nature of the current international segmentation methods limits their usefulness.

In this paper, we apply a new dynamic segmentation methodology, based on semiparametric modeling, to six new product categories in 79 countries and show that, in this sample, country segmentation in new product growth is dynamic and not static. Our approach makes it possible to identify markets that are homogeneous during a given part of the product life cycle. We find that country segment membership varies over the product life cycle and that accounting for this time variation provides superior prelaunch forecasts. Therefore, we recommend that international firms and public policy bodies (e.g., the European Commission and the United Nations) that have a stake in understanding cross-national differences in innovativeness and in making global forecasting reports reconsider their current practice. These entities should adopt a dynamic segmentation approach instead of an exogenously defined regional segmentation approach or a static model-based segmentation approach. In so doing, they should also cautiously consider the set of products or product categories to use as a reference (i.e., the estimation sample) when deriving dynamic segments. This choice is likely to affect the outcome of the segmentation and should therefore be made with

Table 5 Model fit comparison.

	Number of segments	BIC	Log-likelihood
<i>Parametric models</i>			
Single-country segments	79	56,198	-21,680
A priori-defined segments (geographic regions)	8	48,658	-23,643
Static segments	4	44,601	-21,926
Dynamic segments	2	44,712	-22,060
<i>Semiparametric models</i>			
Single-country segments	79	70,211	-11,808
A priori-defined segments (geographic regions)	8	41,118	-18,866
Static segments	4	36,331	-17,125
Dynamic segments	3	36,187	-16,883

Table 6

Prelaunch mean absolute forecast errors (MAD) per product in the hold-out sample, for various prediction horizons from $h = 1$ to $h = 5$ for all models^a.

DVD Players					
	$h = 1$	$h = 2$	$h = 3$	$h = 4$	$h = 5$
<i>Parametric models</i>					
Single-country segments	0.78	1.14	1.49	1.83	2.15
A priori-defined segments	0.52	0.85	1.30	1.80	2.45
Static segments	0.43	0.70	1.20	1.76	2.55
Dynamic segments	0.37	0.71	1.22	1.77	2.44
<i>Semiparametric models</i>					
Single-country segments	0.68	1.00	1.32	1.68	2.04
A priori-defined segments	0.27	0.45	0.75	1.07	1.48
Static segments	0.28	0.45	0.72	1.05	1.49
Dynamic segments	0.15	0.23	0.32	0.48	0.64
Internet Subscribers					
	$h = 1$	$h = 2$	$h = 3$	$h = 4$	$h = 5$
<i>Parametric models</i>					
Single-country segments	0.80	1.11	1.41	1.87	2.58
A priori-defined segments	0.00	1.35	2.27	3.12	4.03
Static segments	0.00	1.58	2.72	3.72	4.75
Dynamic segments	0.00	1.16	1.86	2.43	3.06
<i>Semiparametric models</i>					
Single-country segments	0.68	1.00	1.25	1.62	2.22
A priori-defined segments	0.00	0.65	1.11	1.66	2.42
Static segments	0.00	0.66	1.19	1.84	2.68
Dynamic segments	0.00	0.33	0.69	1.31	2.22

^a The lowest mean absolute deviations (MAD) are given in bold.

care. In our data, we find a substantial amount of heterogeneity across products, which we captured by using the product-deviation function.

Our research can be extended in multiple ways. First, our dynamic segmentation is done at the aggregate level, and our method identifies countries, rather than consumers, that share similar penetration patterns. Country-level analysis conveys a number of advantages, such as the excellent availability of data at the country level and the good accessibility and cost-effectiveness achieved through centralization of the resulting country segments. However, country segments also suffer from a number of limitations. They overlook the differences that exist between consumers within these countries and ignore the potential horizontal consumer segments that cross national borders. Country segments also tend to be less responsive to marketing efforts than disaggregated consumer segments (Steenkamp & Ter Hofstede, 2002). An interesting avenue for future international segmentation research is to collect individual (adoption) data for an international sample of consumers for cross-national dynamic segmentation purposes. One could also extend our methodology, as done by Bijmolt et al. (2004), to combine country and consumer segments.

Second, our approach makes an abstraction of the role of country-specific and product-specific characteristics that might partially underlie the diffusion processes. For example, the observed penetration level in one country may be influenced by the adoption of a product in neighboring countries (e.g., cross-country spillover effects and lead-lag effects). Similarly, influences can also occur between products, as could occur in the presence of competitive or substitution effects for multigenerational technologies (Islam & Meade, 2010). While such effects are outside the scope of the present research, it would be a fruitful research goal to study their role in diffusion by including them at two levels of the hidden Markov model: (i) in the response component of the HMM as additional covariates and/or (ii) in the specification of the transition probabilities, in the same manner as proposed by Netzer et al. (2008).

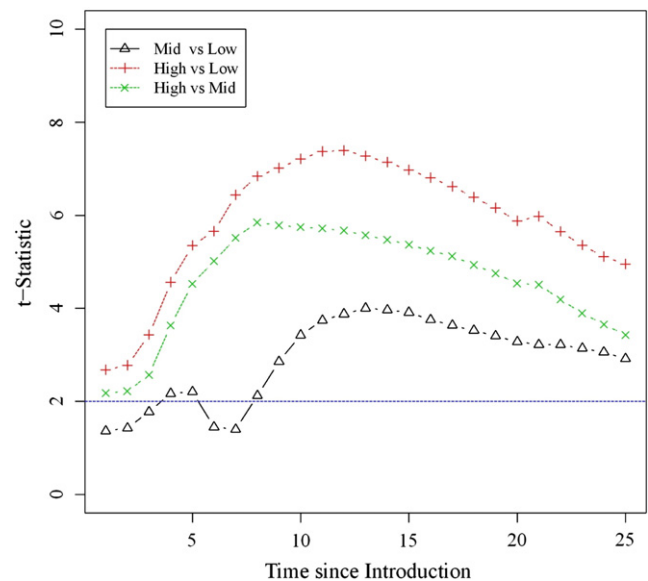
Third, we show that country segments are intrinsically dynamic, but the question remains how the complexity of the segments that these

methods yield can be absorbed by managerial practice. Therefore, more research is needed to understand how organizations could align themselves to have more dynamic international structures.

Overall, our dynamic segmentation framework opens multiple opportunities to tackle nonstationary phenomena in marketing where segmentation is needed. As the increasing number of studies in this area testifies, modeling marketing dynamics will clearly be one of the important research areas in marketing science in the next decade.

Appendix

Value of the t-statistic for testing the equality between every pair of segment-specific curves over time. The horizontal line is the 5% critical value.



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